The role of greasyback shrimp (*Metapenaeus ensis*) in improved extensive shrimp farming systems in the Mekong Delta, Vietnam

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Abstract: In 2010, a variety of factors affecting the size and yield of *Metapenaeus ensis* were measured on shrimp farms in the My Xuyen (MX) and Dong Hai (DH) Districts, Mekong Delta, Vietnam. Studies were conducted in shrimp-rice paddy rotation systems (in MX) and in integrated shrimp-mangrove systems (in DH). A survey was conducted in 120 households, and samples were obtained from a total of 8 farms in the 2 districts. Results showed that *M. ensis* seed was collected from wild populations through tidal recruitment at sluice gates. The recruitment occurred multiple times per year. Yields of *M. ensis* in MX and DH were 31.0 ± 5.9 kg • ha⁻¹ • year⁻¹ and 39.2 ± 4.7 kg • ha⁻¹ • year⁻¹, respectively. The yields were positively correlated with the frequency of water exchange and with the depths of ditches and flat-forms (p < 0.01). The contributions of *M. ensis* to the total shrimp yield in MX and DH were 6.0% and 12.2%, respectively, with respect to yield, and 8.5% and 7.3%, respectively, with respect to value.

Keywords: Greasyback shrimp, Metapenaeus ensis, shrimp-rice paddy rotation systems, integrated shrimp-mangrove systems

1. Introduction

Shrimp farming plays an important role in the societal and economic development of coastal communities in Vietnam (JOHNSTON *et al.*, 2000a; SINH, 2009). The area devoted to shrimp cultivation expanded from 90,000 ha in 1991 to 430,000 ha in 2003 (Loc, 2003), and the yield increased from 55,316 tons in 1995 to 413,132 tons in 2009 (ANONYMOUS, 2010a). Shrimp export values in 2007 were over US\$ 1.2 billion (Tu *et al.*, 2008).

In Vietnam, reported shrimp yields are mainly for the black tiger shrimp (*Penaeus*

* Corresponding author : Tel/Fax : 81-3-5463-0563 Email : sakurak@kaiyodai.ac.jp *monodon*), because this is the only species for which fishery managers have collected data and conducted annual assessments. The species has been studied by authors from both Vietnam and abroad from many viewpoints, such as fishery techniques, market factors, and societal and economic development. However, P. monodon is not the sole commercial shrimp species in Vietnam. Many other high-value species (e.g., Metapenaeus ensis, Metapenaeus lysianassa, and Penaeus merguiensis) are also harvested on shrimp farms. The seed sources for these (non-P. monodon) species are wild populations, referred to as "natural shrimp," collected by tidal recruitment (JOHNSTON et al., 2000b). In contrast, artificial seed has been used in P. monodon since 1990 (NHUONG and HA, 2005).

The Mekong Delta (MD) in the southern region of Vietnam, with an area of about 39,000 km², contains 70% of the nation's shrimp farming. A variety of farming systems are used in this region.

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In traditional extensive shrimp farming systems, the wild seed is trapped and held until harvest. Stocking densities are low, food supplements are not provided, and there is no fertilization (BRENAN *et al.*, 2006). In improved extensive shrimp farming systems, the traditional system is modified by supplementation with artificial stocks of *P. monodon* (stocking density, 1-7 individuals • m⁻²), either with or without feeding (PHUONG *et al.*, 2006; ANONYMOUS, 2007). Finally, in semi-intensive and intensive systems, which have steadily expanded since 2000 (VIET, 2006), *P. monodon* is cultured using intensive food and habitat enrichment schemes to produce rich harvests.

Metapenaeus ensis, which is caught and cultured as part of the traditional shrimp farming activities of local residents in Southeast Asia (LING, 1973), tolerates a wide range of salinity (5-30) and is in high local market demand because of its high meat quality (LIAO and CHAO, 1983; KING, 2001). In the Gulf of Carpentaria, Australia, the maximum total length is 154 mm for males and 189 mm for females (GARCIA, 1985). The longevity of *M. ensis* is approximately 14 months, as found in Hong Kong, China (LEUNG, 1991). Various aspects of this species have been investigated in different countries, such as its reproductive biology in Australia (COURTNEY et al., 1989), its population dynamics in a shrimp pond in Hong Kong (LEUNG, 1997), and the modeling of population dynamics in Osaka Bay, Japan (TAGUCHI et al., 2002).

In this study, we focused on the status of M. ensis in improved extensive shrimp farming systems in Vietnam. These systems constitute about 75% of the total shrimp farming area in MD. The results of this study contribute to the development of a critically needed database on resources vital to the socioeconomic well-being of populations in Southeast Asia, especially in the MD area.

2. Materials and Methods

2.1. Study sites

This study was conducted in 2010 in the My Xuyen (MX) district of Soc Trang Province and in the Dong Hai (DH) district of Bac Lieu Province in the MD (Fig.1). These locations are



Fig. 1. Location map of Vietnam, the Mekong Delta, and sampling sites in the My Xuyen (MX) and Dong Hai (DH) Districts.

representative of the social, economic, and environmental characteristics of the MD region. MD, with its tropical climate, is biseasonal with a rainy season from June to November and a dry season from December to May (H_{UNG}, 2009). An estimated 19% (786,329ha) of the total area is affected by changing water salinity, with predominately fresh water in the rainy season and brackish water in the dry season (VuoNG and LIN, 2001). The most important locations for improved extensive shrimp farming systems are (1) in rice paddies by using shrimp-rice crop rotation and (2) in integrated shrimp-mangrove habitats.

The shrimp-rice paddy rotation system, which was introduced 30-40years ago (VuONG and LIN, 2001; PRESTON *et al.*, 2003), consists of rice cultivation in the wet season (on flat-form areas in predominately fresh-water conditions) and shrimp culturing in the dry season (in predominately saline-water conditions) (BRENNAN *et al.*, 2002). Rice is cultivated from June/July to September/October, when salinity is 0-3, while shrimp are cultured from January/ March to May/June, when salinity is greater than 3.

The total land area of the MX district is



Fig. 2. Construction of farms in the My Xuyen and Dong Hai Districts, designed for improved extensive shrimp culture farming. (a) Relative proportions of ditch and flat-form areas; (b) and (c) Transects of a shrimp-rice paddy rotation system and an integrated shrimp-mangrove system located in MX and DH, respectively, showing (1) ditch depth; (2) depth of flat-form; (3) ditch; (4) flat-form; and (5) sluice gate.

54,450ha (ANONYMOUS, 2010b). The area of shrimp-rice paddy rotation increased from 500ha in 1982 to 6,635ha in 1988 (PRESTON et al., 2003) and to 20,200ha in 2010 (ANONYMOUS, 2011a). Combined shrimp-rice farming provides enhanced income and economic benefits to farmers, as compared with that provided by the cultivation of either shrimp or rice alone (VUONG and LIN, 2001). Mangrove forests occupy 52,786ha in the DH district (ANONYMOUS, 2010c), in regions where salinity ranges from 28 to 34 (ANONYMOUS, 2010d). Shrimpmangrove farms are designed so that mangrove trees occupy flat-form areas and shrimp are cultured throughout the year. The operational principles of this system are described by BINH et al. (1997) and JOHNSTON et al. (2000a). The design of these farms is shown in Fig. 2.

2.2. Data collection

Information was collected in a survey of 60 households from each district, it contains farming area, depth of flat-form and ditch, frequency of water change, yields and prices of M. ensis and P. monodon and other shrimps, investment and income. The participants were farmers who were applying the improved extensive system in these districts.

Table 1. Areas and abbreviations of farms in the two districts where samples were collected.

| My X | uyen | Dong Hai | | |
|---------------|----------------|---------------|--------------|--|
| Farm | Area (ha) | Farm | Area (ha) | |
| AMX | 1 | ADH | 4.3 | |
| BMX | 1.5 | BDH | 2.5 | |
| CMX | 1 | CDH | 3 | |
| DMX | 0.6 | DDH | 3.4 | |
| Mean \pm SD | 1.03 ± 0.4 | Mean \pm SD | $3.3\pm~0.8$ | |

Shrimp sampling was conducted at a total of 8 farms in the MX and DH districts (4 farms in each district). Table 1 shows the area of each farm. The distance between farms in both districts was about 0.2 km. Shrimp culturing/harvesting in MX occurs only during the dry season; therefore, sampling was conducted only between January and June. Metapenaeus ensis in MX farms was collected at night by using trap nets, as shown in Fig. 3 (a), with a mesh size of 1.5 - 1.7 cm. The trap nets were submerged in the ditch, as shown in Fig. 2 (b-3). Sampling was performed on 2 nights, for approximately 8h each night (9:00PM to 5:00AM), by using 6 trap nets at each farm during each sampling time.

In contrast, shrimp culturing in DH occurs



Fig. 3. Trap net design used to catch shrimp in ditches (a) and sluice gates (b).

throughout the year, and sampling was conducted in February, April, June, August, October, and December in 2010. The sampling was conducted during the spring tide of each month, drainage was controlled using a sluice gate, as shown in Fig. 3 (b), through which shrimp were recruited and harvested. *Metapenaeus ensis* was sampled on 2 nights, for 3-4h per night, by using a single sluice gate at each farm.

Shrimp samples collected inside the farms included shrimp obtained via new recruitment and old shrimp from previous recruitments. The samples of M. ensis were kept in ice prior to the measurement of carapace length (CL). CL was measured to the nearest millimeter by using calipers and defined as the shortest distance between the posterior margin of the orbit and the mid-dorsal posterior edge of the carapace. In addition, salinity was measured as an important physical characteristic of the water (MACIA, 2004).

2.3. Data analysis

The non-parametric Wilcox test was used to test the significance of differences between the observed variables in the 2 districts. Akaike's Information Criterion (AIC) (AKAIKE, 1973) was used to select the variables that affected the yield, and these variables were used in the multivariate regression models (DALGAARD, 2002; FARAWAY, 2005).

In this case, the dependent variable, y, is defined as the yield of *M. ensis*, and the maximum number of independent variables, which were defined later, is 6. The number of models, which are constructed using a single independent variable $(x_1, x_2, \dots, \text{ or } x_6)$, can be calculated by the combination C(6, 1) = 6, i.e., $y = f_1(x_1)$, y $= f_2(x_2), \dots, y = f_6(x_6)$. The number of models, which are constructed using 2 independent variables (x_1, x_2) , (x_1, x_3) , ..., or (x_5, x_6) , can be calculated by the combination C(6, 2) = 15, i.e., $y=f_7(x_1, x_2), y=f_8(x_1, x_3), \dots, y=f_{21}(x_5, x_6)$. The total number of models can be calculated by the summation of the combinations $\sum_{i=1}^{b} C(6, i) = 64.$ We calculated the AIC values for each of the 64 models and selected the optimal model as the one with the lowest AIC (FARAWAY, 2005). When the linear regression equation is expressed as follows:

$$\hat{y}=\hat{lpha}\!+\!\hat{eta}_1x_1\!+\!\hat{eta}_2x_2\!+\!\cdots\!+\!\hat{eta}_kx_k$$

with k+1 parameters $(\hat{\alpha}, \hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k)$, and the

| Table 2. Parameters of the improved extensive shrimp farming systems in the MX (mainly shrim) | р |
|---|---|
| -rice paddy rotation systems) and the DH (mainly integrated shrimp-mangrove fores | t |
| systems) (mean±standard deviation). | |

| Item | My Xuyen (n=60) | Dong Hai (n=60) |
|---|---------------------------|---|
| Period of shrimp culture (month•year ⁻¹) | 6 | 12 |
| Farming area (ha•household ⁻¹) | $0.9{\pm}0.3^{a}$ | $1.7 {\pm} 0.6^{\text{b}}$ |
| Depth of flat-form (m) | $0.5 {\pm} 0.3^{a}$ | $0.4{\pm}0.1^{	ext{b}}$ |
| Depth of ditch (m) | $0.9{\pm}0.1^{a}$ | $0.8{\pm}0.1^{	ext{b}}$ |
| Frequency of water exchange (times•year ⁻¹) | $6.6 {\pm} 1.2^{\circ}$ | $19.5 \pm 1.8^{\text{b}}$ |
| Yield of M. ensis (kg •ha ⁻¹ •year ⁻¹) | $31.0 {\pm} 5.9^{\circ}$ | $39.2 \pm 4.7^{\circ}$ |
| Coefficient of variation of M. ensis yield (CV)(%) | 18.9 | 11.9 |
| Price of M. ensis (US\$•kg ⁻¹) | $3.4{\pm}0.5^{\circ}$ | $3.5 {\pm} 0.1^{\circ}$ |
| Yield of P. monodon (kg•ha ⁻¹ •year ⁻¹) | $539.5 \pm 265.7^{\circ}$ | $322.2 \pm 91.7^{\circ}$ |
| CV of P. monodon yield (%) | 49.3 | 28.4 |
| Price of P. monodon (US\$•kg ⁻¹) | $5.3 {\pm} 0.4^{a}$ | $6.7 {\pm} 0.7^{	ext{b}}$ |
| Yield of aquatic other species (kg•ha ⁻¹ •year ⁻¹) | $30.5 \pm .5.7^{\circ}$ | $29.2 {\pm} 7.6^{\circ}$ |
| CV of aquatic other species (%) | 18.8 | 26.0 |
| Price of aquatic other species (US\$•kg ⁻¹) | 1.7 ± 0.2^{a} | $1.5{\pm}0.1^{	ext{b}}$ |
| Total investment from shrimp $(US\$\cdot ha^{-1}\cdot year^{-1})$ | $861.5 \pm 179.7^{\circ}$ | $308.5 {\pm} 71.6^{\circ}$ |
| Net income from shrimp $(US\$\cdotha^{-1}\cdot year^{-1})$ | $2,380 \pm 1,666^{\circ}$ | $2,\!108\!\pm\!741.8^{\scriptscriptstyle \mathrm{b}}$ |
| CV of income from shrimp (%) | 70 | 35.1 |
| Total cost from rice paddy US\$•ha ⁻¹ •crop ⁻¹) | 934 ± 130 | 0 |
| Net income from rice paddy $(US\$ \cdot ha^{-1} \cdot crop^{-1})$ | 645.7 ± 117.6 | 0 |
| Coefficient of variation of rice paddy (%) | 18 | Invalid |
| Total of net income $(US\$\cdot ha^{-1}\cdot year^{-1})$ | $2,444.8\pm1,650^{\circ}$ | $2,108\pm741.8^{	ext{b}}$ |
| Total of net income $(US\$\cdothousehold^{-1}\cdot year^{-1})$ | $2,716.5\pm1,485^{\circ}$ | $3,\!584\!\pm\!1,\!681^{	ext{b}}$ |
| CV of net income household $^{-1}$ year $^{-1}$ (%) | 54.6 | 36.3 |

Mean values of parameters in the same row with the different superscripts show significant different $(p \le 0.01)$

residual sum of squares is expressed as $RSS = \sum_{i=1}^{n} (\hat{y}_i - y_i)^2$, then AIC is calculated by :

$$AIC = \log\left(\frac{RSS}{n}\right) + \frac{2k}{n}$$

where n is the number of samples.

The density function was used to determine the variations in CL for each sampling time. The CL had a normal distribution; analysis of variation and Tukey's honestly significant difference test were used to test for the significance of CL variations between different months.

To relate these data to characteristics of the shrimp culture system, 8 variables (defined later) were selected from the surveys for analysis. These variables were considered to have effects on the system in terms of the aquaculture (MACIA, 2004), and they were hypothesized to have effects on the yield and CL of M. ensis. Factor analysis (FA) was used to reduce the original number of variables and identify new variables that explained the most important variances in the data.

3. Results

3.1. Contribution of *M. ensis* to households

The contribution of *M. ensis* to the total shrimp yield in the MX and DH districts in 2009 was 6.0% and 12.2%, respectively, in terms of physical yield, and 8.5% and 7.3%, respectively, in terms of crop value. The yields of

| | My X | Kuyen | Dong Hai | |
|--|----------|----------|----------|----------|
| | Factor 1 | Factor 2 | Factor 1 | Factor 2 |
| <i>P. monodon</i> yield $(kg \cdot ha^{-1} \cdot year^{-1})$ | 0.9 | | 0.99 | |
| Depth of flat-form (m) | | | 0.77 | |
| Depth of ditch (m) | 0.50 | | 0.64 | |
| Net income $(US\$ \cdot ha^{-1} \cdot year^{-1})$ | 0.97 | | 0.95 | |
| <i>M. ensis</i> yield $(kg \cdot ha^{-1} \cdot year^{-1})$ | | 0.87 | | 0.94 |
| Frequency of water exchange (times•year ⁻¹) | | 0.70 | | 0.79 |
| Yield of other species $(kg \cdot ha^{-1} \cdot year^{-1})$ | | | | |
| Total cost of shrimp production (US \bullet ha ⁻¹ •year ⁻¹) | | | | |
| Coefficient of determination | 0.26 | 0.16 | 0.37 | 0.21 |
| Total coefficient of determination | | 0.42 | | 0.58 |

Table 3. Results of factor analysis, based on eight parameters obtained from the survey of 120 households in My Xuyen and Dong Hai. The eight parameters were extracted from Table 2.

M. ensis were found to be more stable than those of *P. monodon*. The coefficients of variation for *M. ensis* yields were 18.9% and 11.9% in the MX and DH districts, respectively, while those for *P. monodon* yields in the 2 districts were 49.3% and 28.4%, respectively (Table 2).

According to farmers in MX, the main purpose of food supplements is to feed *P. monodon*, because this species is stocked at a density of 5-6 individuals • m⁻². In contrast, in DH the land area per household is larger than the land area of MX households (p < 0.01) (Table 2), and *P. monodon* is stocked at lower densities $(1-3 \text{ individuals } \cdot \text{m}^{-2}) 3-5$ times per year. This pattern of stocking is referred to as thinning (by harvesting) and subsequent stock compensation, enabling shrimp farms to maintain a low density of individuals without food supplements.

The total investment for shrimp culturing in MX was higher than that in DH (p < 0.01). However, the situation reversed with respect to the net income from shrimp culturing in the 2 districts, which was significantly different (p < 0.01) (Table 2). However, farmers in MX also obtain an income from rice paddy harvests in the rainy season (Table 2) ; therefore, their income per hectare is actually higher than that of the DH residents. While DH is located in the ecological mangrove region, the region can not grow rice paddy. Besides, shrimp culture activities of farmers in the region are used depending much on natural environment, low investment and low yield, farming area of each household in DH is larger area than in MX, thus, their income per household are higher than in MX (Table 2).

FA was used to identify 2 orthogonal linear combinations of the 8 original parameters measured in MX (Table 3). These 2 factors are shown in Fig.4 (a). Factor 1 has 3 components, representing a positive correlation between the depth of the ditch and the yield of *P. monodon* and net income. Factor2 has 2 components, representing a positive correlation with the number of water exchanges and the yield of *M. ensis*. These factors explained 42% of the original variance; factors 1 and 2 explained 26% and 16%, respectively, of the variance in the data.

Similarly, the first 2 factors of the FA in DH explain 58% of the variance in the 8 original parameters measured in DH (Table 3). These factors (Fig. 4 (b)) showed that factor 1 has 4 components, representing a positive correlation among the yield of *P. monodon*, the depth of the flat-form, the depth of the ditch, and the net household income. Factor2 has 2 components, representing the yield of *M. ensis* and the number of water exchanges. Factors 1 and 2 explained 37% and 21%, respectively, of the variance in the DH data.



Fig. 4. Plot of the first 2 factor analysis loading vectors for My Xuyen (a) and Dong Hai (b). Me : M. ensis yield; we : number of water exchanges; Ot : yield of other species; fd : depth of ditch; ffw : depth of flat-form; tc : total cost; ni : net income; and Pm : P. monodon yield.



Fig. 5. Correlation between frequecies of water exchange and *M. ensis* yields in My Xuyen (a) and Dong Hai (b).

3.2. Variables affecting the yield of *M. ensis*

Farmers in MX often add more water to their farms to compensate for leakage or evaporation, and they limit the drainage of water away from farms because brackish water is scarce and the quality is variable. We found that the number of water exchanges in DH was higher than that in MX (p < 0.01; Table 2) and that yields of *M. ensis* were positively correlated with the number of water exchanges (Fig. 5).

Ditches and flat-forms were deeper in MX than in DH (p < 0.01). Results of the multivariate regression analysis were as follows. When 6

independent variables $(x_{1sr}, x_{2sr}, x_{3sr}, x_{4sr}, x_{5sr}, and x_{6sr})$ were used, the model explained 46% of the variance in the yield of *M. ensis*, where $x_{1sr}, x_{2sr}, x_{3sr}, x_{4sr}, x_{5sr}, and x_{6sr}$ denote the rate of water exchange, depth of ditch, depth of flat-form, yield of other species, yield of *P. monodon*, and price of *P. monodon*, respectively. The number of water exchanges (x_{1sr}) was significant in all the cases; this variable is a significant predictor of *M. ensis* yields (p < 0.01). The optimal model that minimized the AIC value (AIC=184, R² = 0.45) was as follows:

$$y_{sr} = 3.6 + 3.18 x_{Isr} + 0.96 x_{2sr} + 11.35 x_{3sr}$$
(1)

Table 4. Predictions for optimal yields of *M. ensis* in the My Xuyen (MX) and Dong Hai (DH) districts with multivariate regression models. Subscriptions of *sr* and *sm* denote the shrimp rice in MX and shrimp mangrove in DH, respectively. *y* denotes the *M. ensis* yield. x_1 , x_2 , x_3 , x_6 , x_5 , and x_6 denote the frequency of water exchange, depth of ditch, depth of flat-form, yield of other species, yield of P. monodon, and quantity of *P. monodon* stocked, respectively.

| MX (N=60) | | | DH (N=60) | | | |
|---|-----|----------------|--|-----|----------------|--|
| Model | AIC | \mathbb{R}^2 | Model | AIC | \mathbb{R}^2 | |
| $y_{sr} = f(x_{1sr}, x_{2sr}, x_{3sr}, x_{4sr}, x_{5sr}, x_{6sr})$ | 188 | 0.46 | $y_{sm} = f(x_{1sm}, x_{2sm}, x_{3sm}, x_{4sm}, x_{5sm}, x_{6sm})$ | 144 | 0.59 | |
| $y_{sr}{=}f\left(x_{1sr},\ x_{2sr},\ x_{3sr},\ x_{4sr},\ x_{5sr} ight)$ | 186 | 0.46 | $y_{sm} = f(x_{1sm}, x_{2sm}, x_{3sm}, x_{4sm}, x_{5sm})$ | 143 | 0.58 | |
| $y_{sr} = f(x_{1sr}, x_{2sr}, x_{3sr}, x_{4sr})$ | 185 | 0.46 | $y_{sm} = f(x_{1sm}, x_{2sm}, x_{3sm}, x_{4sm})$ | 141 | 0.58 | |
| $y_{sr} = f(x_{1sr}, x_{2sr}, x_{3sr})$ | 184 | 0.45 | $y_{sm} = f(x_{1sm}, x_{2sm}, x_{3sm})$ | 140 | 0.57 | |
| | | | $y_{sm} = f(x_{1sm}, x_{2sm})$ | 139 | 0.56 | |
| | | | $y_{sm} = f(x_{1sm})$ | 138 | 0.56 | |



Fig. 6. Correlation between yields of *M. ensis* and depth of flat-form in My Xuyen (a) and Dong Hai (b).

This optimal model can explain 45% of the variance in *M. ensis* yields (Table 4).

Similarly, in DH, the model constructed using 6 independent variables (x_{1sm} , x_{2sm} , x_{3sm} , x_{4sm} , x_{5sm} , and x_{6sm}) explained 59% of the variance in M. ensis yields. The optimal model (AIC=138, $R^2=0.56$) was constructed using only one variable, x1sm, and explained 56% of the variance in M. ensis yields. The model is represented as follows :

$$y_{sm} = 1.167 + 1.950 \ x_{Ism} \tag{2}$$

The results of both univariate (Fig. 5) and multivariate regression analyses (Table 4) confirmed that water exchange is the principal factor that enhanced yields of M. ensis. In addition, these results also showed that farms with deeper flat-forms had higher yields of M. *ensis* than those with shallow flat-forms (Fig. 6). Furthermore, yields of M. *ensis* were higher in farms with deeper ditches (p < 0.01) in both shrimp-rice paddy rotation and integrated shrimp-mangrove systems.

3.3. Variations in CL

Size density diagrams of *M. ensis* in MX are shown in Fig. 7, with CL for the early crop in January with a peak concentration of 18.0 -23.0 mm. Other months, however, show broader (less concentrated) density peaks for CL. CL appeared to vary between 18.0 and 30.0 mm, and there were significant differences in CL between the different months (p < 0.01),



Fig. 7. Size density of carapace length (CL) for M. ensis in the My Xuyen District at each of the 6 sampling times; n: total number of shrimp in 4 farms per month.



Fig. 8. Size density of carapace length (CL) for M. ensis in the Dong Hai District at each of the 6 sampling times; n: total number of shrimp in 4 farms.

except for February-March (p>0.05). Similarly, the size density structure of CL for the 6 sampling periods in DH (Fig. 8) showed significant differences in CL between the different months (p<0.01), except for June-August (p>0.05).

Differences in salinity, depth of ditch, and depth of flat form for farms in MX and DH are shown in Tables 5 and 6. The flat-form depth of the farms and depth of the ditches in both MX and DH did not differ with respect to location or the months sampled. However, the salinity values were quite different between the locations. The salinity values in DH were much higher than those in MX. Furthermore, in MX, the salinities in April–June were higher than those in January–February.

4. Discussion

Coastal shrimp farming started in Vietnam in the 1970s, with development of extensive farming systems. *Metapenaeus ensis* farming has declined in recent years, possibly because of mangrove forest clearing, over-harvesting, and urbanization (JOHNSTON *et al.*, 2000a). The natural shrimp harvest in MX was 181 kg •ha⁻¹

| | January (N=16) | February (N=16) | March (N=16) | April (N=16) | May (N=16) | June (N=16) |
|---------------------|-------------------|--------------------|-----------------|-----------------|---------------|----------------|
| Flat-form depth (m) | $0.4{\pm}0.1$ | $0.4{\pm}0.1$ | $0.45{\pm}0.1$ | $0.5{\pm}0.1$ | $0.5{\pm}0.1$ | $0.5{\pm}0.1$ |
| Ditch depth (m) | $0.8{\pm}0.2$ | $0.7{\pm}0.2$ | $0.8{\pm}0.2$ | $0.8{\pm}0.2$ | $0.9{\pm}0.1$ | $0.8{\pm}0.2$ |
| Salinity | $15.5 {\pm} 2.0$ | 15.5 ± 1.5 | 15 ± 2.2 | 19.7 ± 2 | 20.0 ± 3 | $19.5{\pm}2.3$ |

Table 5. Depth of ditches (m), depth of flat-forms (m) and salinity measured during the six sampling periods in MX (mean±standard deviation).

Table 6. Depth of ditches (m), depth of flat-forms (m) and salinity measured during the six sampling periods in DH (mean±standard deviation).

| | February (N=16) | April (N=16) | June (N=16) | August (N=16) | October (N=16) | December (N=16) |
|---------------------|--------------------|-----------------|-----------------|------------------|-------------------|--------------------|
| Flat-form depth (m) | $0.4{\pm}0.2$ | $0.4{\pm}0.2$ | $0.4 {\pm} 0.1$ | $0.45{\pm}0.2$ | $0.4 {\pm} 0.2$ | $0.4 {\pm} 0.2$ |
| Ditch depth (m) | $0.8{\pm}0.2$ | $0.7{\pm}0.5$ | $0.8 {\pm} 0.1$ | $0.8{\pm}0.1$ | $0.9{\pm}0.2$ | $0.8{\pm}0.2$ |
| Salinity | $25.0 {\pm} 2.5$ | 30.0 ± 2 | 31.0 ± 1.2 | 30.0 ± 1.2 | 29.0 ± 3.1 | 29.0 ± 2.2 |

in 1997 (BE et al., 2003). Currently, the harvest is about 31.0 ± 5.9 kg·ha⁻¹·year⁻¹. The harvest of *M. ensis* in mangrove habitats was 100-600kg·ha⁻¹ in 1993-1994 (BINH and LIN, 1995), whereas now it is about 39.2 ± 4.7 kg·ha⁻¹· year⁻¹. Thus, the income generated by *M. ensis* harvests is declining. This species, however, is still very important to local residents because natural stocks are free. The farming of *M. ensis* has been a substantial help to poor shrimp farmers in the MD, where the GDP per capita was US\$ 1,113 in 2009 (ANONYMOUS, 2011b).

Shrimp culturing is the primary income source for residents in both MX and DH districts, with combined *M. ensis* and *P. monodon* polycultures common in all farms. However, variations in *P. monodon* yields probably reflect household investment levels. For this reason, the *P. monodon* yields were selected as one of the main components of factor 1 and could explain the variability in the *M. ensis* yields.

Multivariate analyses confirmed that M. ensis and P. monodon populations on shrimp farms are not in competition. The correlation between M. ensis and P. monodon yields in the MX and DH districts was significantly low, i.e., 0.02 and 0.11, respectively. This is because P. monodon was stocked at a low density in these systems, and the wild-stock density of M. ensis was also low.

Water exchange is the best mechanism for recruitment of post-larval *M. ensis* in farms. We found that *M. ensis* yields depend strongly on the number of water exchanges. However, the waters of the MD coastal region are highly turbid (ANONYMOUS, 2010e). Sediment can quickly fill ditches and accumulate on flatforms, especially if water exchanges occur multiple times in a year. Farmers must then incur the costs of annually excavating the sediment (PRESTON *et al.*, 2003). Yields of *M. ensis* are also dependent on factors such as wild-seed abundance, the water levels of flooded farms, farm design, and water quality (BINH *et al.*, 1997).

Our results show that recruitment of M. ensis is strongly dependent on the frequency and volume of water exchanges on farms. The size density structures of M. ensis populations varied throughout the year, with one to several peaks in the CL per month; CL varied between 18 mm and 30 mm. CL varied from month to month in both MX and DH, implying that yields of this species are sustained throughout the year.

The yield of M. ensis in MX was lower than that in DH because M. ensis in MX is affected to a greater extent by natural conditions. Shrimp culture farming operates only during the dry season (6 months) in MX. However, local residents can get other income from cultivating paddy rice in the same area. It is an effective way to use agricultural land, and it can be considered to be applicable to coastal areas. The integrated shrimp-mangrove system also has low yields. However, the average area per household is 1.7 ha. The existence of integrated shrimp-mangrove farms not only provides a livelihood for local residents but also helps to sustain mangrove forests. Both of these systems require low investments and are a consistent income source for local residents.

Although *M. ensis* yields are lower than those of *P. monodon*, the species contributes significantly to the alleviation of poverty in coastal communities (JOFFRE and SCHMITT, 2010). *Metapenaeus ensis* not only supplements the incomes of farmers but also has positive impacts on regional biodiversity and contributes to the natural balance of ecological systems (ISLAM *et al.*, 2004)

5. Conclusions

Populations of *M. ensis* are combined with those of *P. monodon* in improved extensive throughout farming systems MD. Metapenaues ensis populations rely on the recruitment of wild seed. with multiple recruitments during the year representing the opportunity for sustained yields. Metapenaeus ensis cultures yield significant economic support to low-income populations, requiring only small investments to realize modest returns. The shrimp are harvested for the market; however, yields of M. ensis are lower than those of P. monodon. For this reason, farmers cannot depend on *M. ensis* alone for their livelihood.

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Application of chemical mass balance to determination of phytoplankton composition from pigment profiles in seawater

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Abstract: We applied a chemical mass balance (CMB) receptor model to determine the composition of phytoplankton from pigment profiles in seawater collected from Sendai Bay, Japan. In descending order, the relative abundances of the six predominant phytoplankton classes expressed as a proportion of total chlorophyll *a* were Bacillariophyceae > Prasinophyceae \geq Cryptophyceae \geq Haptophyceae \geq Pelagophyceae \geq Dinophyceae. The results obtained by using CMB model were similar to those obtained by the CHEMTAX software: the Bacillariophyceae accounted for about 65% (CMB) or 64% (CHEMTAX) of the total phytoplankton, and the proportions of the other classes ranged from 5% to 11%. Many DNA sequences obtained from the clone library of the *psbA* gene encoding the D1 protein were highly similar to those of the Bacillariophyceae. The DNA sequencing results were similar to those obtained by CMB. We anticipate that CMB, which is used to investigate sources of pollutants, can be applied as a simple method of evaluating phytoplankton community composition using a pigment ratio approach.

Keywords: phytoplankton pigment, chemical mass balance analysis, DNA sequencing.

1. Introduction

Phytoplankton are primary producers in the ocean and use many kinds of pigments for photosynthesis. Pigment profiles vary among phytoplankton classes. Analysis of pigment profiles in seawater has been used to investigate seasonal changes in phytoplankton dominance (METAXATOS and IGNATIADES 2002, RODRIGUEZ *et al.* 2003, HASHIHAMA *et al.* 2008), enabling higher throughput than can be achieved by cell counting under a microscope. Phytoplankton pigments are analyzed by using high-performance liquid chromatography (HPLC); the proportional concentration of each phytoplankton class is calculated statistically by, for example, factor analysis using a steepest descent algorithm (CHEMTAX; MACKEY et al. 1996) or multiple regression using ordinary least squares analysis (BUSTILLOS-GUZMAN et al. 1995, SUZUKI et al. 1997, 2002). CHEMTAX software (MACKEY et al. 1996, 1997) is frequently used for the calculations (Suzuki et al. 2002, Rodriguez et al. 2003. HASHIHAMA et al. 2008) and CHEMTAX v. 1 is an excellent free software package, but it functions only within a commercial programming language.

Chemical mass balance (CMB) software is used to solve mass balance equations by an effective variance least-squares method. It was developed by Dr. TOM COULTER of the U.S. Environmental Protection Agency and was

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configured by HAYAKARI and HANAISHI (2001) to operate with add-in Microsoft Excel software (Microsoft, U.S.A.). Because Microsoft Excel is the world's most popular spreadsheet software, CMB has the advantage of ease of use because it does not need a special programming language. Moreover, CMB is different from previous multiple regression analyses of phytoplankton pigments, as it is based on a mass balance equation with a weighted leastsquares regression analysis (HOPKE 2002). The CMB receptor model is used in environmental studies to investigate sources of pollutants, such as airborne particulate matter in the atmosphere (HOPKE 2002)and dioxins (KASHIWAGI et al. 2006, OKUMURA et al. 2008).

If the datasets of mass concentrations of particulate matter components are replaced by pigment profiles, CMB, like CHEMTAX, can be used to estimate phytoplankton composition by a pigment ratio approach. As CMB software (ver. 8J0612) can be used when only few data are available, we expected it to be suitable for monitoring phytoplankton communities in coastal areas in cases where there are only a few sampling points. Here, we used the CMB receptor model to estimate phytoplankton community composition from pigment profiles in seawater. We then compared the results of CMB analysis with those of CHEMTAX and DNA sequencing, both of which have recently been used to investigate plankton diversity (SAVIN et al. 2004, ZEIDNER and BEJA 2004, ZHU et al. 2005, JING et al. 2010).

2. Materials and methods

2.1. Sampling and HPLC analysis of pigments

In May 2007, a seawater sample was collected from the ocean surface off Oginohama in Sendai Bay, Japan (lat $38^{\circ}23'$ N, long $141^{\circ}24'$ E). This area contains fertile grounds used for Pacific oyster (*Crassostrea gigas*) aquaculture. The 200-ml sample was filtered through a glass fiber filter (25-mm Whatman GF/F filter, Whatman PLC, Springfield Mill, U.K.). The pigments were extracted from the filter for 24 h in 1 ml methanol, and the extract was centrifuged at $10000 \times g$ for 10 min. After dilution of the supernatant to 80% with distilled water in an HPLC auto-sampler (Shimadzu, Kyoto, Japan), the pigments were analyzed by the method of ZAPATA *et al.* (2000), with a slight modification whereby a guard column was attached between the injection valve and the analytical column.

Each pigment was identified from the retention times of pigment standards, and the concentration was calculated from the peak area on the chromatograph from the HPLC diode array detector. We used eight pigment standards: 19'-butanoyloxyfucoxanthin (Butfuco). 19'-hexanoyloxyfucoxanthin (Hexfuco), chlorophyll a (Chl-a), chlorophyll b(Chl-b), fucoxanthin (Fuco), neoxanthin (Neo), and peridinin (Perid) were purchased from Wako Chemicals (Osaka, Japan) and DHI (Hørsholm, Denmark), Laboratory and alloxanthin (Allo) was refined by HPLC from pigment extracts of unialgal cultures of Rhodomonas salina CS-174 (Cryptophyceae), purchased from CSIRO Marine Research (Hobart, Tasmania, Australia). The concentrations of standard pigments dissolved in acetone or ethanol were calculated from the extinction coefficients ($E_{1 cm}$; L g⁻¹ cm⁻¹), as measured by spectrophotometer (GE Healthcare Ultrospec 3000) over a 1-cm cuvette path length. The $E_{1 cm}$ and measurement wavelength, and the solvent used for each pigment, were obtained from published data (JGOFS 1994, JEFFREY et al. 1997). The concentrations of pigment standards were calculated as follows (JGOFS 1994):

$$C_{\rm s} = (A_{\rm max} - A_{750\,\rm nm})/E \times 1000 [\rm mg/g]$$
 (1)

where $C_{\rm s}$ is the pigment concentration (mg L⁻¹), $A_{\rm max}$ is the absorbance maximum, $A_{750 \,\rm nm}$ is the absorbance at 750nm to correct for light scattering, and E is the extinction coefficient (L g⁻¹ cm⁻¹).

2.2. Analysis by CMB and CHEMTAX

The pigment concentrations determined by HPLC were input into a CMB spreadsheet in Microsoft Excel. Data for the typical pigment ratios of each class of phytoplankton were also input (Table 1). We selected six phytoplankton — *Phaeodactylum tricornutum* (Bacillariophyceae), *Pelagococcus subviridis* (Pelagophyceae), *Pycnococcus provasolii* (Prasino-

| | Perid | But-fuco | Fuco | Neo | Hex-fuco | Allo | Chl b | Chl a |
|-------------------|-------|----------|-------|-------|----------|-------|---------|-------|
| Prasinophyceae | | | | 0.151 | | | 0.945 | 1.000 |
| Bacillariophyceae | | | 0.755 | | | | | 1.000 |
| Dinophyceae | 1.063 | | | | | | | 1.000 |
| Haptophyceae | | | | | 1.706 | | | 1.000 |
| Pelagophyceae | | 0.368 | 0.974 | | | | | 1.000 |
| Cryptophyceae | | | | | | 0.229 | | 1.000 |

Table 1. Typical pigment / chlorophyll a ratios of each class of phytoplankton used in this study

phyceae), Amphidinium carterae (Dinophyceae), Emiliania huxleyi (Haptophyceae), and Chroomonas salina (Cryptophyceae) and calculated their pigment ratios from the published data (JEFFREY and WRIGHT 1997) used for the initial ratio matrix in the CHEMTAX user's manual (MACKEY et al. 1997). After inputting all data, we analyzed the phytoplankton profiles in the environmental sample by CMB according to the CMB user's manual (HAYAKARI and HANAISHI 2006).

The mass environmental concentration of pigment i (C_i) is shown in Equation (2) : (in the original equation, C_i is the mass concentration of component i of particulate matter);

$$C_{i} = \sum_{j=1}^{p} a_{ij} S_{j} \quad i = 1,, n$$
 (2)

where a_{ij} is the fractional concentration of pigment *i* of class *j* (in the original equation, a_{ij} is the density of particulate matter containing component *i* in emissions from source *j* to the receptor); S_j is the total mass concentration contributed by class *j* (originally the total mass concentration contributed by source *j*); *p* is the number of classes and *n* is the number of pigments, with $n \ge p$ (originally, *p* was the number of sources and *n* was the number of components). C_i and a_{ij} are known, and S_j is found from the effective variance least-squares solution of an over-determined system of equations.

The calculations are accompanied by an error in the fractional concentration of the pigment from each class (or, in the Air Pollution Control Technology Manual [ENVIRONMENT AGENCY, GOVERNMENT of JAPAN 1997], the emission source profile data) and by an error in the established data on the environmental concentration of the pigment (or, in the abovementioned manual, the concentration of the component on environmental particles). Both of these errors must be considered. The effective variance least-squares method is solved in Equation (2) by multiplying the concentration of each pigment (or, in the manual, the concentration of each chemical) by a percentage weight that is proportional to the fractional concentration of the pigment from each class (or, in the manual, the precision of the emission source profile data) and the measured environmental concentration of each pigment (or, in the manual, the environmental concentration of each chemical component). Repeat computations are conducted to seek S_i , which minimizes the function x^2 in Equation (3):

$$x^{2} = \sum_{j=1}^{n} \frac{(C_{i} - \sum_{j=1}^{p} a_{ij}S_{j})^{2}}{V_{i}}$$
(3)

In Equation (4) V_i is the effective variance :

$$V_{i} = \sigma c_{i}^{2} + \sum_{j=1}^{p} \sigma_{aij}^{2} S_{j}^{2}$$
(4)

where σ_{ci} is the error that accompanies component *i*, and σa_{ij} is the error that accompanies the measurement of a_{ij} .

To evaluate whether the calculated values were appropriate, the χ^2 , R^2 , and percent mass values were also calculated by CMB. In the CMB manual, the χ^2 value is the weighted sumof-squares of the differences between the calculated and measured fitting species concentrations; a value less than 1 indicates a very good fit to the data, and values between 1 and 2 are acceptable. The R^2 value is the fraction of the variance in the measured concentrations that is explained by the variance in the calculated species concentrations. R^2 ranges from 0 to 1.0; the closer the value is to 1.0, the better the source contribution estimates explain the measured concentrations. Percent mass is the ratio of the model-calculated source contribution to the measured mass concentrations, expressed as a percentage. This ratio should equal 100%, although values ranging from 80% to 120% are acceptable (COULTER 2004, HAYAKARI and HANAISHI 2006).

After the data for the typical pigment ratios of each phytoplankton class were input (Table 1), we also used CHEMTAX to analyze the phytoplankton community profile in the same environmental sample according to the CHEMTAX user's manual (MACKEY *et al.* 1997).

2.3. DNA sequencing and BLAST similarity search

For DNA sequencing, a 500-ml seawater sample was filtered through a Durapore membrane filter (diameter, 47 mm; pore size, $0.45 \,\mu$ m; Millipore, Merck KGaA, Darmstadt, Germany). The filter containing the sample was suspended in 1 ml of lysis buffer (20 mM Tris • HCl [pH8.0], 5 mM EDTA, 0.3% (wt : vol) SDS, 200 μ g/ml proteinase K). The suspension was vortexed for 30s and then incubated at 55°C for 1 h and at 95℃ for 5 min. The filter was then removed from the lysis buffer and the liquid phase containing nucleic acids was separated by centrifugation at $10,000 \times g$ for 10 min. The supernatant diluted with distilled water was used as a template for polymerase chain reaction (PCR) analysis.

The *psbA* gene, which encodes polypeptide D1 of the photosystem II reaction center complex, was amplified by PCR with the degenerate primers PsbAF (5'-TTC GGT CAA GAA GAA GAG ACT TA-3') and PsbAMR (5'-YTC RTG CAT HAC YTC VAW RCC-3'), which have specificity for many algae. Primers were selected by using Primer3 software (ROZEN and SKALETSKY 2000). PCR was performed under the following conditions: 94°C for 10 min; 40 cycles of 94°C for 30s, 55°C for 1 min, and 72°C for 1 min; and a final step at 72°C for 7 min.

After PCR amplification, the PCR products

were purified with a High Pure PCR Cleanup Micro Kit (ROCHE DIAGNOSTICS, Germany). The purified products were cloned by DYNA-EXPRESS TA PCR Cloning Kit (BIODYNAMICS LABORATORY INC., Tokyo, Japan). After colony-direct PCR (30 cycles of 98°C for 10 s, then annealing and extension at 68°C for 1 min) with the universal primers M13P7 (5'-CGC CAG GGT TTT CCC AGT CAC GAC-3') and M13P8 (5'-AGC GGA TAA CAA TTT CAC ACA GGA AAC-3'), colonies containing inserts were identified by agarose gel electrophoresis. After purification of the PCR products of the insert colonies using ExoSAP-IT (GE HEALTHCARE, England), DNA sequencing was performed with a DYEnamic ET Terminator Cycle Sequencing Kit (GE HEALTHCARE) by using the primer M13P8 on an automated ABI 3100 DNA sequencer (APPLIED BIOSYSTEMS USA). Sequences were then aligned by using phred/phrap/consed software (Ewing et al. 1998, GORDON et al. 1998, SAITO 2009) and phytoplankton species were identified from sequences aligned by using BLAST.

3. Results and discussion

3.1. Measured pigment concentrations and concentrations calculated by the CMB model

We compared measured and calculated pigment concentrations (Fig. 1). Measured pigment concentrations were in the order Chl-a >



Fig. 1. Comparison of measured and calculated pigment concentrations in a seawater sample from Sendai Bay, Japan. Error bars were calculated by assuming a 5% error in the measured concentrations.



Fig. 2. Relative contributions of six phytoplankton classes to total chlorophyll a in a sample of seawater from Sendai Bay, Japan, as calculated by CMB and CHEMTAX.

Fuco > Chl-b. Hex-fuco, Perid, Allo, But-fuco, and Neo were at the same low level as Chl-b. The calculated concentration of each pigment was consistent with the measured concentration. The χ^2 , R^2 , and percent mass values calculated by CMB were 0.62, 0.99, and 101.6%, respectively. We consider that these three values indicate the validity of the class contribution estimates.

Chl-a, which is in all phytoplankton, usually has the highest concentration of all pigments detected in environmental samples (WRIGHT et al. 1991, ZAPATA et al. 2000), as it did in this study. Fucoxanthin, which is found in the Bacillariophyceae and Pelagophyceae (JEFFREY and WRIGHT 1997), was also detected at high concentrations. Our pigment profiles are therefore similar to those reported elsewhere (WRIGHT et al. 1991, ZAPATA et al. 2000).

3.2. Comparison between CMB and CHEMTAX

We compared the relative contributions of six phytoplankton classes to total chlorophyll *a* calculated by CMB and CHEMTAX (Fig. 2). Both approaches indicated that the relative abundance of the Bacillariophyceae was highest, followed in descending order by Prasinophyceae \geq Cryptophyceae \geq Haptophyceae \geq Pelagophyceae \geq Dinophyceae. The results of the CMB and CHEMTAX analyses were similar. The Bacillariophyceae accounted for about 65% (CMB) or 64% (CHEMTAX) of all phytoplankton (Fig. 2). The proportions of Prasinophyceae, Cryptophyceae, Haptophyceae, Pelagophyceae, and Dinophyceae were about 11%, 8%, 6%, 5%, and 5% (both CMB and CHEMTAX), respectively. In Sendai Bay, the Bacillariophyceae outnumber the Dinophyceae in all seasons (ILJIMA *et al.* 2004) ; therefore, we believe that the phytoplankton community composition determined in this study is characteristic of that in Sendai Bay.

The pigment concentrations in phytoplankton are known to vary with environmental conditions such as light and nutrient conditions, and among species or strains within the same class (MACKEY *et al.* 1998, SCHLUTER *et al.* 2000, LIONARD *et al.* 2008). As the change of the pigment : chlorophyll a ratios of phytoplankton accompanying environmental conditions are reflected in the CMB analyses, we anticipate that CMB is a useful tool for investigating phytoplankton composition in the environment.

3.3. DNA sequencing and BLAST similarity search

The BLAST search showed that many of the *psbA* gene DNA sequences were highly similar to those of the Bacillariophyceae (Table 2).

| Phylogenetic group | Sequences producing significant alignment (GenBank accession no.) | Identity (%) | Number of aligned reads | Ratio of each phylogenetic group (%) |
|--------------------|--|-----------------|----------------------------|--|
| Bacillariophytceae | Odontella sinensis (Z67753) | 268/285 (94%) | 12 | 63.6 |
| | Thalassiosira pseudonana (EF067921) | 256/260 (98%) | 2 | |
| Haptophyceae | Isochrysis sp. (AY119753) | 271/282 (96%) | 1 | 13.6 |
| | Emiliania huxleyi (AY741371) | 243/255 (95%) | 1 | |
| | Phaeocystis antarctica (AY119756) | 232/241 (96%) | 1 | |
| Prasinophyceae | Ostreococcus sp. (EU851961) | 272/278 (97%) | 1 | 9.1 |
| | Micromonas pusilla (FJ858269) | 278/285 (97%) | 1 | |
| Rhizaria | Paulinella chromatophora (DQ789030) | 184/215 (85%) | 3 | 13.6 |

Table 2. Results of a BLAST search of DNA partial sequences from water samples from Sendai Bay

Bacillariophyceae sequences accounted for about 64% of all phytoplankton. This proportion is close to the values obtained by CMB (65%) and CHEMTAX (64%). Within the Bacillariophyceae, Odontella sinensis was significantly aligned by BLAST. Odontella spp. are frequently observed in the coastal zones of the Tohoku region of Honshu, Japan (TAKANO 1990); our results are consistent with these observations. The results of the BLAST search for similar DNA sequences suggest that the Haptophyte species were Isochrysis sp., Emiliania huxleyi and Phaeocystis antarctica, and the Prasinophytes were Ostreococcus sp., Micromonas pusilla and Rhizaria was Paulinella chromatophora. Although we did not detect any Cryptophyceae, Pelagophyceae or Dinophyceae DNA sequences, the pigments characteristic of these phylogenetic groups indicate that they accounted for 5.0% to 8.4% of the total phytoplankton, as determined by CMB and CHEMTAX. If more DNA sequences are available, it should be possible to detect the DNA sequences of these classes. Although the bias of PCR frequently adds uncertainty to the quantification of organisms (DIEZ 2001), our DNA sequencing results were roughly similar to those of CMB and CHEMTAX. As more results of CMB analysis accumulate, CMB will become an easy-to-use tool for investigating natural phytoplankton community composition.

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串本・浦神間水位差に現れる季節的変化

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Seasonal variation of the sea level difference between Kushimoto and Uragami tide–gauge stations

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Abstract : Sea level difference between Kushimoto and Uragami tide-gauge stations has been used as an indicator of flow pattern in the sea south of Japan: the difference is small when the Kuroshio has meandering path and is large when it has straight path. The difference exhibits seasonal variation, and it becomes large in the summer season in the years when the Kuroshio takes straight path throughout year as in 2003 and in 2009. We analyzed the sea level data of Kushimoto and Uragami tide-gauge stations for 17years from 1994 through 2010. It is shown that the Kuroshio takes straight path when the separation distance of the northern edge of the Kuroshio is smaller than 15km from the tip of Cape Shionomisaki. The sea level difference data were selected only for the time that the separation distance is smaller than 15km, and the average sea level for 17 years are calculated. Significant seasonal variation can be seen in the averaged sea level difference curve. However, the seasonal variation is not so clear in comparison with cases in 2003 and 2009. The sea level changes at Kushimoto and Uragami tide-gauge stations were investigated for the years of 2003 and 2009. Both of the sea levels tend to rise in summer time, but the magnitude of the rise at Kushimoto is considerably larger than that at Uragami. MAEKAWA et al. (2011) showed that the water off the Kushimoto tide-gauge station is originated to the surface water of the current zone of the Kuroshio, when the Kuroshio takes straight pass.. Thus, the main cause of seasonal variation of the sea level difference would be sought for seasonal warming of the surface water of the Kuroshio area. The water would be brought from upstream (southern) area of the Kuroshio, and would be heated much more than surface water in Kumano-nada in summer season.

Keywords : Sea level difference between Kushimoto and Uragami, seasonal variation, intrusion of the Kuroshio Water off the southwest coast of Kii Peninsula, water off Kumano –nada

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Fig. 1 The distance (upper figure in km) between the northern edge of the Kuroshio and the tip of Cape Shionomisaki, and the sea level difference (lower figure in cm) between Kushimoto and Uragami tide gauge stations in 2003. The sea level differences are shown with solid circles when the separation distance is shorter than 15km, and with open circles when the separation distance is longer than 15km.

1. はじめに

われわれは、紀伊半島南西海岸に見られる振り 分け潮をはじめとして, 紀伊半島の先端付近の微 細海況の解析を行ってきた(TAKEUCHI et al., 1998. NAGATA et al., 1999, UCHIDA et al., 2000, 中村ら, 2008, 田中ら, 2008, 前川ら, 2011, 2012)。その中で、前川ら(2011)は、串本・浦 神検潮所間の水位差に季節変化があり、水位差が 夏期に大きくなる可能性があることを指摘してい る。このような傾向は 2003 年や 2009 年など年間 を通して黒潮が直進路を維持している年に明確に 現れる。しかし、顕著な季節変化が見られない年 も少なくない。和歌山県農林水産総合技術センター (以後和歌山水試と略記)では、潮岬灯台近くに 目視観測点を設けて,毎日午前9時前後に,沖合 数百 m の海面を観測し,潮波の立ち方の目視観 測を行っており、漁海況情報の一つとして活用し ている。目視観測者は夏期に潮波が大きくなり, 黒潮の接岸流が強くなる傾向があると指摘してい る(このことは、別に検討する予定である)。

前川ら(2011)は、黒潮流路の指標となる串本・ 浦神の水位差は、潮岬半島沖の東西わずか数 km の部分で生じていることを指摘すると共に、水位 差を生じさせているのは、300m 以浅の海洋構造 であることを示した(このことは藤田(2001)が 指摘しているが、その詳細な水平分布を提示した のは初めてである)。このような表層の構造は、 黒潮強流帯域においても、明確な季節変動を示し ており、振り分け潮に伴って紀伊半島南西岸にも たらされる黒潮系水の特性も季節変化を示すであ ろう。振り分け潮の発生そのものは、福田ら (2002)が数値試シミュレーションで再現に成功 しているが、このような表層構造や微細海況構造 に関しては、観測による資料の収集が、現在でも 必要とされる。本論文では、1994 年から 2010 年 までの 17 年間の資料を解析するとともに、顕著 な季節変化を示す 2003 年および 2009 年の資料を 解析し、この季節変化の特性を検討する。

2. 串本・浦神間水位差の季節変化

2-1 年々の季節変化

前川ら(2011)は串本・浦神間水位差の季節変 化について、その Fig. 2 に 2009 年の変化を示し ている。しかし、黒潮流路の変遷もあり、前述し たような典型的な変化が毎年見られるとは限らな い。ここでは、2009 年と同様に明確な季節変化 が見られた 2003 年について、串本・浦神間水位 差の変化を Fig. 1 の下段に示す。図の上段に衛



Fig. 2 Same as in Fig.1 except for in 2004.

星資料から求めた黒潮北縁の潮岬先端からの距離 (以下では簡単のため黒潮離岸距離と呼ぶ)の変 化を示してある。これは、和歌山水試が NOAA/ HRPT 受信解析装置(TeraScan)を用いて直接 受信している。衛星の画像の分解能は、1.1km 程度とされているが、実際の画像からはより細か く黒潮北縁の離岸距離を読みとることが可能であ り, ここでは 0.1km 単位で決めた値を用いてい る。ただし、陸と海のコントラストが大きいこと もあり、黒潮北縁が潮岬先端から約3.7km 以内 に近づくと北縁位置の決定は難しくなり、この場 合は黒潮離岸距離を3.7kmとした。また、海洋 保安庁海洋情報部が発行している海洋速報では黒 潮流軸の位置が記載されており、これは黒潮北縁 より24km沖側とされている。しかし、流軸位置 が24kmより岸寄りに示されている場合があり, この定義に従えば、北縁の位置は陸上にあること になる。海洋速報の情報を参考にしているが、衛 星資料を基にした北縁位置を本論文で論じている。 水位差については串本・浦神のそれぞれの観測値 を東京湾中等潮位(TP)上に換算した後、その 差をとり、水位差0が、両水位が同レベルに対応 するように表記している。

前川ら(2011)が示した2009年と同様に, 2003年は年間を通して黒潮は安定した直進路を とっており、水位差が夏期に高くなるという季節 変化が、明瞭に現れている。参考のためFig.2 に 2004 年の水位差の変化を同様に黒潮離岸距離 と共に示すが、この年は、黒潮の流路が7月に潮 **岬から離れて、それ以後は典型的な蛇行路をとる** ようになった年である。7月以降の水位差は5cm 以下の安定した値をとるようになるとともに、季 節変化は全く見られなくなった。Fig. 1 および Fig. 2 では、黒潮離岸距離が潮岬先端から 15km 以内の場合を黒丸で、15km以上の場合を白丸で 示してあるが、これらの図から 15km 以内の場合 のみをとると、黒潮が蛇行路をとっている場合の データをほぼ除くことができることがわかる。な お、ここでも黒潮の直進路、蛇行路という表現を 便宜上用いているが、少なくとも潮岬近傍の海況 に影響するのは、日本南岸沖の全体的な流路特性 ではなく,黒潮北縁が潮岬からどれだけ離れてい るかの離岸距離である。離岸距離は、潮岬の東西 の沿岸水が一体的なものであるか、両者が分離さ れているかを決定する要因である。

Fig. 2 に見られるように黒潮が蛇行路をとり, 黒潮北縁の位置が潮岬先端から遠く離れると, 串 本・浦神間の水位差は非常に小さくなる。以下に 1994 年から 2010 年までの 17 年間の資料の解析 を行うが, このように短い期間の解析においては, その期間内で蛇行路をとっていた頻度の高い月の 水位差の平均値は小さくなり, 直進路をとってい た頻度の高い月の水位差の平均値は大きくなる。 われわれの解析期間において, 蛇行路をとってい



Fig. 3 The sea level difference between Kushimoto and Uragami when the separation distance of the northern edge of the Kuroshio is less than 15km from the tip of Shionomisaki from 1994 to 2010.



Fig. 4 Same as in Fig. 3 except for the daily average for the 17 years. The thick curve indicates 11 days running mean.

た頻度は、2~4月の期間が、他の月に比べてか なり大きく、2~4月の水位差は見掛け上小さく なってしまう。そこで、以下の解析では、黒潮離 岸距離が15km以内の場合のみを選んで行うこと にする。

黒潮離岸距離が 15km 以内の場合のみを選び出 して、1994 年から 2010 年までの 17 年間のそれ ぞれの年の水位差の変化を示したのが Fig. 3 で ある。ただし、ここでは、見やすくするため閏年 については 12 月 31 日のデータを除いてある。こ の図を見ると、季節変化がほとんど見られない年 もあるが、全体的に見て 8~9 月頃を中心として 水位差が大きくなっている傾向が認められる。黒 潮離岸距離 15km 以下の場合のみの各日付につい ての平均値を示したのが Fig. 4 である。直進路 のデータのみを用いているから、水位差は年間を



Fig. 5 Variation of the sea level difference averaged for each ten days: (upper figure) for the case that the separation distance of the northern edge of the Kuroshio is less than 15km from the tip of Shionomisaki, (middle figure) for the case the separation distance is from 15km through 30km, (lower figure) for the case the separation distance is larger than 30.1km. The vertical lines attached to data points indicate error bars.

通して正の値を示している(正の値は串本側の水 位が浦神側の水位よりも高いことを表す)。解析 期間が短いため1日毎の平均値の曲線には細かい 振動が現れる。11日間の移動平均をとった曲線 もFig.4に示してあるが、この移動平均によっ て、細かい振動はほぼ除かれている。いずれにせ よ、8~9月頃を中心として水位差が大きくなる 季節変化は平均値の曲線に明確に現れている。11 日間の移動平均をとることによって、スムーズな 曲線が得られたことから、以下の検討では、1日 毎の平均ではなく、その11日間の移動平均値、 または旬平均値を基に議論することにする。

2-2 旬平均水位差の季節変動と黒潮離岸距離

旬平均を求めるに当たり、大の月では、1~10 日を上旬、11~20日を中旬、21~31日を下旬と 定義し、2月を除く小の月では、これに対して下 旬を21~30日と定義した。2月については1~10 日を上旬とするのは変わりないが、11~19日を 中旬とし,20日から28日,あるいは29日まで を下旬と定義した。この定義に従って,串本・浦 神間の水位差の旬平均を計算して示したのが Fig.5である。1994年から2010年までの17年 間のデータから,黒潮離岸距離を15.0km以内 (上図),15.1~30.0km(中図),30.1km以上(下 図)の3つの場合について示した。ただし,解析 期間についてのデータ数は、3つの場合について, 互いにかなり違っており,15.0km以内の生起日 数が57.3%,15.1~30.0kmのそれが23.4%, 30.1km以上のそれが19.3%である。

Fig. 5 には各計算値に対して標準偏差σを縦棒 で示してあるが, Fig. 3 に見られるように年々変 動が大きいため, 標準偏差の値はかなり大きくなっ ている。これらの平均値の曲線を見ると, 15.0km 以内の場合にのみ, 9 月頃をピークとし て, 夏期に水位差が大きくなる季節変化を認める ことができる。この結果は, 夏期に水位差が大き くなる現象は, 黒潮が直進路をとり, 潮岬の東西



Fig. 6 Variations of the sea level at Kushimoto (solid circle) and at Uragami (open circle) are shown in upper figure in 2003. Sea level difference between two stations is shown in lower figure.



Fig. 7 Same as in Fig 6 except for in 2009.

の沿岸水が分離された場合にのみ,水位差に季節 変動が現れることを示唆している。より明確な結 論を得るには,解析期間をより長くとり,データ 数を増す必要があるが,過去に解析期間を延ばす ことは,衛星データを含め,データの取得や,利 用できるデータの精度等に問題がある。この論文 では,明確な季節変化が見られた2003年と2009 年の場合を検討することによって,季節変化の生 じる原因を考えることにする。

3. 2003 年及び 2009 年に見られた串本, 浦神の それぞれの水位の季節変化

2003 年と 2009 年の串本および浦神の水位の変 化を Fig. 6 及び Fig. 7 にそれぞれ示す。これら の値は、東京湾中等潮位(TP)からの変位とし て表し、1時間間隔の水位データにタイドキラー フィルター(花輪・三寺、1985)を施した後、毎 正午の値を抜き出して示したものである。また、 これらの図の下段には両検潮所間の水位差を再掲 してある。

2003年,2009年共に、串本・浦神の両地点で、 夏期の水位が上昇する傾向を示す。これは大気か らの加熱による表層水温の上昇に伴う現象として 解釈されよう。しかし、夏期の水位上昇は、串本 の上昇量が浦神のそれよりもはるかに大きく、そ れによって両地点間の水位差が形作られているこ とがわかる。

従って,水位差の季節変化を起こす主要因は,

串本側に求めざるを得ない。黒潮離岸距離が 15km以内であるときは、通常黒潮は直進路をとっ ており、振り分け潮の発生に伴い黒潮流域の表層 水が紀伊半島南西岸に侵入してくる時である。主 温度躍層より下の黒潮水が顕著な季節変化を示さ ないが、主温度躍層以浅の黒潮表層水の水温は季 節変化を起こし、大気からの加熱により夏期上昇 する。紀伊半島南西岸に侵入してくる黒潮表層水 の水温が夏期には上昇し軽くなり、串本側の水位 の夏期の上昇を起こすことは当然のことと考えら れる。もちろん、正確な上昇量を求めるには塩分 の変化も論ずる必要がある。

Fig. 6 および Fig. 7 に見られるように, 浦神 の水位も若干夏期に上昇する。南方から運ばれて くる黒潮表層水の夏期の水温上昇は、浦神沖の沿 岸水よりも大きいことは十分考えられる。しかし、 潮岬西方で振り分け潮に伴って沿岸域に侵入して くる水が、黒潮のどの部分の水であるかが明らか にされていない以上、定量的な検討は出来ない。 われわれは、水の動きを明確にするため、水塊分 析や種々のトレーサー解析を進めているが十分に 成功するに至っていない。前川ら(2012)は、溶 存酸素の鉛直プロファイルの微細構造に注目した 解析を試みて、潮岬東方の黒潮流の中に熊野灘沖 の沿岸水が黒潮流域に取り込まれ、運び去られる ことを示唆している。もしそれが正しければ、浦 神沖の沿岸水が常に更新されていて、そのため季 節変化が抑制される可能性がある。このような沿 岸水の更新は、黒潮が長期にわたって直進路を維 持している場合,効果的であろう。今まで,われ われは、振り分け潮の関連で潮岬西方の海況につ いて詳細な検討を進めてきた。しかし、串本・浦 神間の水位差の季節変化のような問題を論じるた めには、熊野灘沖の海況特性についても詳細な検 討を行う必要があろう。

4. おわりに

串本・浦神間の水位差が夏期に増大するという 季節変化が現れることは、黒潮直進時に紀伊半島 南西岸に振り分け潮が現われさらに、その南西海 岸に侵入する黒潮水は、黒潮域の表層の水である とする前川ら(2011)の推論を支持するものであ る。しかし、ここで論じてきたように、この水位 差の季節変化を起こす主要因は、紀伊半島南西岸 に侵入する黒潮表層水の季節変化であるが、熊野 灘側の沿岸水の季節変化が、この水位差の減少に 寄与していることも忘れてはならない。振り分け 潮の発生についても、今後解明すべき問題が山積 している。福田ら(2002)は、数値実験で振り分 け潮の発生を再現することに成功しているが、こ れは、南西沿岸域にもたらされる水が表層水に限 られ、振り分け潮の発生が水深 100m 程度の位置 で起こっているような、微細な構造を説明するも のではない。前川ら(2011)が論じているように、 振り分け潮の発生に結びつくような水平水温・塩 分場の等値線の岸向きの膨らみが観測されている ものの、ADCP によって観測された流速場には その部分で岸に向かう流れは見出せない。振り分 け潮の発生のメカニズムは、まだ十分に解明され ていない。これを明らかにするには、黒潮流域の どこの水が、どのような経路に沿って沿岸域にも たらせられるのかを正確に知ることが必要である。

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マタナゴ当歳魚におけるアマモ場からアラメ場への成長に伴う移動

吉永 潔

Ontogenetic habitat shift in yearling surfperch Ditrema temmincki pacificum from a seagrass bed to a seaweed bed at Tateyama, central Japan.

Kiyoshi Yoshinaga

Abstract : To confirm ontogenetic migration in yearling surfperch *Ditrema temmincki* pacificum from a seagrass bed to a seaweed bed, fish censuses were carried out in both beds using SCUBA from February to November in 2002 to 2004 at Tateyama, Chiba Prefecture, central Japan. Newborn fish were most numerous in the seagrass bed in May, yearling density tending to be greater in the seaweed bed than in the former in August. A mark-release experiment conducted in the seagrass bed in May and June, 2005, in which yearlings were marked by clipping the operculum before release to the capture site, indicated that individuals were still present at the site a month later. However, marked fish appeared in the seaweed bed in July to September, being absent from the seagrass bed in August and September, demonstrating that *D. temmincki pacificum* yearlings moved from the seagrass bed to the seaweed bed with growth.

Keywords : Ditrema temmincki pacificum, habitat shift, seagrass bed, seaweed bed

1. はじめに

マタナゴDitrema temmincki pacificumは関 東地方から瀬戸内海にかけての太平洋沿岸に分布 するウミタナゴ科魚類の一種である (KATAFUCHI and NAKABO, 2007)。Abe (1969) によれば、本種(原著ではウミタナゴDitrema temminckiとなっているが、KATAFUCHI and NAKABO (2007) に従いマタナゴとした)は生活 史の初期段階にはアマモ場に分布するが、成長に 伴い岩礁域へ生息場を変えることが推察されてい る。このような成長に伴う生息場間の移動は多く の水産有用種で知られており(例えば、

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GILLANDERS et al., 2003; 佐野ら, 2008), 複数 の生息場をセットで維持することが生産の維持に とって重要であると考えられている(小路, 2009)。 ウミタナゴ科魚類は水産有用種であり(櫻井ら, 2008), 生息場の変化の知見は重要である。しか し, 実際にマタナゴがアマモ場からいつ, どのよ うな成長段階で移動するのかについては, 直接的 にまだ確認されていない。

そこで本研究では、マタナゴ当歳魚の生息場間 の移動を明らかにするために、まず、アマモ場と その沖合いにある岩礁域(アラメ場)において当 歳魚の個体数密度の季節変化を潜水観察で調べた。 生息場間の移動があるならば、当歳魚は最初アマ モ場で多く、その後季節の進行に伴い岩礁域で増 加すると考えられる。さらに、アマモ場に出現し た当歳魚に対し、標識・放流調査を実施すること により、生息場間の移動が実際に行われているか どうかを確認した。



Fig. 1. Map of the study site at Hasama, Tateyama, showing the seagrass and seaweed beds in which underwater observations were conducted. \times , collection and release site for yearlings in the mark-release experiment.

2. 調査場所と方法

本研究は千葉県の南端に位置する館山市波左間 (34°58'N, 139°46'E 測地系 WGS84) で実施した (Fig. 1)。同調査地では、防波堤をはさみ西側の 沖合いは岩礁域でアラメ Eisenia bicyclis が,東 側は砂地でアマモ Zostera marina が密生してい る。このアマモ場およびアラメ場においてマタナ ゴ当歳魚がどの月に多く出現するかを明らかにす るために、2002年から2004年の2,5,8および 11月にスキューバ潜水による目視観察を行った。 各月において、それぞれの藻場の水深 5m 以浅に、 1×50mのトランセクトを5本,10m以上離して ランダムに設置し、トランセクト内に出現したマ タナゴ当歳魚の個体数を計数・記録した。トラン セクト1本あたりの観察時間は10分とし、9:00 から15:00までの満潮時に観察を行った。なお, 松清(1963)は山口県瀬戸内海のマタナゴについ て, 全長は満1年で約150mmになると報告してい る。そこで本研究では、全長 150mm 未満の個体を 当歳魚とした。

マタナゴ当歳魚の個体数について、年ごとに月 と生息場間での違いを二元配置分散分析によって 検討した。月間と生息場間で有意な差が存在した 場合には、Scheffé test を用いて、各月間あるい は各生息場間での有意差の有無を検定した。二元 配置分散分析において、月と生息場に交互作用が 存在した場合、各月に対する生息場間の差と各生 息場における月間の差をそれぞれ Scheffé test によって調べた(UNDERWOOD、1997)。なお、本 研究の主目的は、マタナゴ当歳魚の個体数が月ご と、あるいは生息場ごとでどのように異なり、ま た、それらの異なりが 2002 年から 2004 年にわたっ て、同じように認められるかどうかを明らかにす ることである。このため、個体数における年間の 違いについては解析を行わなかった。分散分析に よる検定を行う際には、等分散性および変量の正 規性を得るために、 個体数を対数変換[log (x+1)]した。

マタナゴ当歳魚における生息場の移動を実証す るために、標識・放流調査を行った。2005年5 月28日から6月19日にかけて,合計157個体 (全長 70~90mm)の当歳魚をアマモ場でタモ網等 により採集し、両側の鰓蓋後縁部を数 🔤 切除し た後、採集地点と同じアマモ場へ放流した。この 標識方法はオイカワZacco platypusですでに用い られており、実績がある(福岡県水産海洋技術セ ンター, 2000)。放流後, 6月25日から9月23 日まで、アマモ場とアラメ場においてスキューバ 潜水による目視観察を行った。目視観察では約 24m/分の速度で10分間遊泳し、進行方向の左右 1m ずつの範囲内(トランセクトの面積は2× 240m)に出現したマタナゴの個体数と鰓蓋の切 れ込みの有無を記録した。この観察は調査期間中, 1~3週間の間隔で計9回行い、1回の観察ではア マモ場とアラメ場でそれぞれ3本のトランセクト を設置した。

なお、本研究の実施後、日本産ウミタナゴ属魚 類は KATAFUCHI and NAKABO (2007) によって 2 種 2 亜種に分類された。本研究で観察したほとん どの個体は体側が銀色で背側が暗青色もしくは暗 赤色であり、この特徴は関東地方から瀬戸内海の 太平洋沿岸に分布するマタナゴに一致した。また、 調査地から採集した5個体の標本を同定したとこ ろ、すべてマタナゴであった。したがって、本研



Fig. 2. Mean number of *Ditrema temminckii pacificum* yearlings observed per transect $(50m^2, n = 5)$ in the seagrass and seaweed beds in each month from 2002 to 2004. Vertical bars indicate standard deviation. Asterisked horizontal bars indicate significant differences in yearling densities (*p < 0.05, **p < 0.01). \mathbf{V} , not observed.

究で観察した個体はマタナゴとした。

3. 結果および考察

3.1 個体数の季節変化

アマモ場とアラメ場におけるマタナゴ当歳魚の 個体数の季節変化を Fig. 2 に示した。1 トランセ クトあたりの平均個体数について、年ごとに月お よび生息場間で比較したところ、いずれの年にお いても交互作用が認められた(2002年はF3.32= 10.6, p < 0.0001; 2003 年は $F_{2,24} = 9.79$, p =0.0008; 2004 年は $F_{3,32}$ =19.2, p < 0.0001)。そこ で各月に対する生息場間の差と各生息場における 月間の差をそれぞれ比較した結果、どの年におい てもアマモ場では5月に多くみられ、11月に少 なかった(2002年と2003年はp < 0.01, 2004年 ltp < 0.05)。一方,アラメ場では8月に多く,5 月と11月に少なかった(2002年はp<0.05, 2003年と2004年はp<0.01)。生息場間では、 いずれの年でも5月においてアラメ場よりもアマ モ場で多かった(2002年と2003年はp<0.01, 2004 年は*p* < 0.05)。また, 8 月において 2002 年 と 2003 年は生息場間でほとんど違いがみられな かったものの、2004年ではアラメ場で有意に多 かった $(p=0.0001)_{\circ}$

HAYASE and TANAKA (1980a) によれば,マ タナゴは5月にアマモ場で産仔する。また,産出 直後の稚魚の全長は約50~60mmであると報告さ れている(松清, 1963; ABE, 1969)。本調査地 では,5月に多くの当歳魚がアマモ場で観察され ており,それらの全長は約50mmであった。した がって,マタナゴは本調査地でもこの頃にアマモ 場で産仔していると考えられた。一方,8月にな るとアマモ場の当歳魚の個体数は減少し,アラメ 場で多くなる傾向がみられた。これは当歳魚がア マモ場からアラメ場に移動したことを示唆する。

3.2 標識放流

アマモ場では、放流直後の6月25日と7月3 日に標識個体が観察されたが、その後、まったく 確認されなかった(Table 1)。一方、アラメ場で はその逆の現象がみられ、標識個体は6月25日 には観察されなかったが、その後、7月3日、8 月1日、9月3、11および23日に出現が確認さ れた(Table 1)。標識放流時の全長は70~90mm であったが、9月にアラメ場で確認された標識個 体は約100mmであった。以上の結果は、本調査地 のマタナゴ当歳魚が成長に伴い、アマモ場からア ラメ場へ移動したことを示している。

このような成長に伴う生息場の変化については、 マダイ Pagrus major でも報告されている。この 種の稚魚では、成長とともに増加する摂餌要求量 に対して相対的に餌生物が不足し、新たな生息場 を求めて浅海砂浜域から沖側へ移動すると考えら れている(藤川, 1986)。また、メバル類は春季

Table 1. Mean number (\pm standard deviation, n=3) of marked yearlings per transect (480m²) in the seagrass and seaweed beds in 2005

| Date | Seagrass bed | Seaweed bed |
|--------------|-------------------|-------------------|
| 25 June | $0.67 {\pm} 0.58$ | 0 |
| 3 July | $0.33 {\pm} 0.58$ | $0.33 {\pm} 0.58$ |
| 17 July | 0 | 0 |
| 1 August | 0 | $0.67 {\pm} 0.58$ |
| 20 August | 0 | 0 |
| 28 August | 0 | 0 |
| 3 September | 0 | $0.33 {\pm} 0.58$ |
| 11 September | 0 | $0.33 {\pm} 0.58$ |
| 23 September | 0 | $0.33 {\pm} 0.58$ |
| | | |

にアマモ場で多いワレカラ類を捕食して成長し, 夏季にアマモ場で餌生物が少なくなるとガラモ場 へ移動する(布施, 1962a, b)。産出直後のマタナ ゴはアマモ場でヨコエビ類やワレカラ類などを専 食することが知られている (HAYASE and TANAKA, 1980b)。それら餌生物の現存量はアマ モの消長と同様の周年変化を示し、春に最も多く、 その後は減少して冬季に最も少なくなる(高間, 1980)。また、既往の報告から(吉川, 1978a, b), 夏季におけるヨコエビ類とワレカラ類の1m²あた りの個体数をアマモ場とアラメ場で比較してみる と、両者の個体数密度はアラメ場のほうで多かっ た。これらのことから、本調査地のマタナゴ当歳 魚は産出されたアマモ場においてヨコエビ類など を餌として成育し、その後、餌要求量の増大と、 夏季のアマモ場餌現存量の減少に伴い、餌を求め てアラメ場へ移動したのではないかと考えられた。

謝辞

調査の実施および本論文のとりまとめにあたり ご指導頂いた東京大学大学院農学生命科学研究科 佐野光彦博士,黒倉寿博士,南條楠土博士,島根 大学汽水域研究センター 堀之内正博博士,財団 法人自然環境研究センター 井上隆博士,英文校 閲を頂いた Graham S. Hardy 博士に深謝します。 また調査を実施するにあたり,ご理解とご協力を いただいた波左間漁業協同組合代表理事組合長 佐野錬一氏,佐野秀雄氏,横須賀市自然・人文博 物館館長 林公義氏,および波左間海中公園スタッ フの方々に心よりお礼申し上げます。

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資 料

第50巻第3-4号掲載欧文論文の和文要旨

Tran Van VIET^{L2}, 桜本和美²:ベトナム・メコンデルタ地区における改良型集約的ヨシエビ養殖の役割

ベトナムメコンデルタのミースイエン (MX) 地区とドンハイ (DH) 地区におけるヨシエビ Metapenaeus ensis の養殖生産量に与える要因について分析した. MX 地区はエビ養殖と稲作を交互に行うローテーションシステムを 実施し,DH 地区はマングローブ水域での集約的養殖を実施している.2010年1月から12月までの間,上記2地区 の8 養殖業者から120の漁業者を抽出し,養殖期間や養殖地の面積など22項目について調査した.因子分析法によ り養殖生産量に与える主要因を特定し,その結果をもとに重回帰分析を行った.その結果,(1) 平均生産量は MX 地区で31.0 kg/ha/year,DH 地区で39.2 kg/ha/year,(2) 生産量は養殖池の水の交換頻度と養殖池の深さと高い 相関を示し,(3) 上記2地区におけるエビの全生産量に占める M. ensisの生産量は,MX 地区で6.0%,生産額で 8.5%,DH 地区では生産量で12.2%,生産額で7.3%であった.また,漁業者一人当りの生産額ではDH の方が MX よりも有意に高かった.

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奥村裕¹,神山孝史¹,鈴木敏之²:植物色素を基にした植物プランクトン組成解析へのケミカル・マス・バランス法の適用

仙台湾で採取した海水試料から植物プランクトンの色素組成を分析後、ケミカル・マス・バランス法(CMB)に より、分類群組成の推定を試みた。CMB法による解析の結果、植物プランクトン全体に占める各分類群の割合は、 珪藻>プラシノ藻≧クリプト藻≧ハプト藻≧ペラゴ藻≧渦鞭毛藻の順に少なくなっていた。解析結果を CHEMTAX 法と比較すると、珪藻の割合は CMB 法で 65%、CHEMTAX 法で 64%となり、それ以外の分類群は CMB 法、 CHEMTAX 法ともに 5%~11%となり、CMB 法による解析結果は CHEMTAX 法とおおよそ一致した。また、 PsbA(光化学系 II 反応中心の D1 サブユニットをコードする)遺伝子のクローン・ライブラリから得た遺伝子配列 も、珪藻と相同性が高い配列が多く、CMB 法による解析結果は遺伝子配列の結果ともおおよそ一致していた。CMB 法は化学物質の排出起源の推定以外にも、色素組成を基に植物プランクトンの群集組成を簡易的に評価する方法とし て、利用できる可能性が示された。

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学会記事

- 1.6月10日(日),日仏会館(東京都恵比寿)におい て2012年度総会を開催した。
 - 日 時: 2012 年 6 月 10 日 (日) 15 時 35 分~16 時 10 分
 - 場 所:日仏会館 501 会議室
 - 出席者:61名(出席;13,委任状による出席;48, 定足数;130×1/6=22名)
 - 議 長:小松会長

第1号議案 2011年度事業報告

- 1) 荒川庶務幹事より会員異動状況が報告された。
 2011 年度増減 正会員 +5-7=-2, 学生会員 +3名
 2012 年 3 月末現在 名誉会員 5, 正会員115,
 - 特別会員 10, 学生会員 5, 賛助会員 7
- 2)活動状況 荒川庶務幹事より,評議員会(2011.6.18 日仏 会館),幹事会(2011.5.30 東京海洋大学),総 会(2011.6.18 日仏会館),研究発表会(2011. 6.18 日仏会館)を各1回開催,論文賞2件の 授与,評議員選挙,会長選挙,学会賞委員半数 改選を行ったことについて報告された。 小松会長より,義援金募金活動状況および日仏 共同研究助成申請・採択について報告された。
- 編集関係
 吉田編集委員長より、学会誌「La mer」49巻1/ 2号、3/4号(創立50周年記念号)の発刊について報告された。

各報告ののち,第1号議案は承認された。

第2号議案 2011 度収支決算報告および監査報告

神田会計幹事より,資料1にもとづき2011年度収 支決算が報告された。長島監事より,2011年度の会 計が適切であるとの監査結果が報告された。

各報告ののち、第2号議案は承認された。

第3号議案 2012年度事業案

荒川庶務幹事より、1)総会(1回)、学術研究発表

会(1回), 評議員会(1回), 幹事会(3回)の開催,

2) 学会賞, 論文賞の候補者の推薦と授与, 3) 学会

賞委員選挙(半数改選),について説明された。

吉田編集委員長より,4) La mer 50 巻 1-4 号の発 刊について説明された。

小松会長より、5)東北カキシンポジウム、日仏会 館学術講演会の企画・開催について、説明された。

各項目の説明ののち,第3号議案は承認された。

第4号議案 神田会計幹事より,資料2にもとづき 2012年予算案が説明された。

審議ののち,第4号議案は承認された。

- 第5号議案 荒川庶務幹事より,資料3にもとづき 2012年度-2013年度役員,評議員,学会賞推薦委員 について紹介ののち,第5号議案は承認された。
- 報告事項
 - ①小松会長より、津波で被災した三陸養殖漁業復興のための義援金募金活動の状況が報告された。
 - ②荒川庶務幹事より、アメリカ先進陸水海洋学会 2012年大会を本学会が後援していることについて 報告された。
 - ③内田広報幹事(代理荒川総務幹事)より,学会内 メーリングリストの作成について報告された。

以上

| 収入の部 | | | | |
|--------|-----------|-----------|----------|-----------------------|
| 費目 | 予算額 | 決算額 | 増 減 | 備考 |
| 前年度繰越金 | 736,584 | 736,584 | 0 | |
| 正会員会費 | 1,048,000 | 936,000 | 112,000 | 131 名 (8,000 円×131 名) |
| 特別会員 | 96,000 | 60,000 | 36,000 | 16 名(6,000 円×16 名) |
| 学生会員会費 | 24,000 | 12,000 | 12,000 | 6名(4,000円×6名) |
| 賛助会員会費 | 130,000 | 140,000 | ▲10,000 | 7社(10,000円×13口) |
| 学会誌売上金 | 164,800 | 120,000 | 44,800 | |
| 広告料 | 60,000 | 20,000 | 40,000 | |
| 別刷り代等 | 461,150 | 500,000 | ▲38,850 | 別刷り,超過頁,カラー印刷費 |
| 掲載料 | 150,000 | 800,000 | ▲650,000 | 3 編×50,000 円 |
| 雑収入 | 73,298 | 100,000 | ▲26,702 | 学術著作権使用料他 |
| 寄付金 | 550,000 | 0 | 550,000 | 笹川財団 |
| 収入合計 | 3,493,832 | 3,424,584 | 69,248 | |

平成23年度収支決算

| 支出の部 | | | | |
|----------|-----------|-----------|----------|-----------------|
| 費目 | 予算額 | 決算額 | 増 減 | 備考 |
| 学会誌印刷費 | 1,645,800 | 1,800,000 | ▲154,200 | 49巻1.2号 49巻3.4号 |
| 送料・通信費 | 104,720 | 150,000 | ▲45,280 | |
| 事務費 | 612,635 | 700,000 | ▲87,365 | 人件費,事務用品,封筒他 |
| 交通費 | 11,100 | 20,000 | ▲8,900 | |
| 会議費 | 29,924 | 5,000 | 24,924 | |
| 学会賞経費 | 18,194 | 50,000 | ▲31,806 | 賞状他 |
| 50周年記念事業 | 0 | 0 | 0 | |
| 雑費 | 18,230 | 25,000 | ▲6,770 | 振込み手数料他 |
| 次年度繰越 | 1,053,229 | 674,584 | 378,645 | |
| 支出合計 | 3,493,832 | 3,424,584 | 69,248 | |

平成24年度予算(案)

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| 収入の部 | | | | |
|--------|-----------|-----------|----------|---------------------|
| 費目 | 24 年度予算 | 23 年度予算 | 増減 | 備考 |
| 前年度繰越金 | 1,069,241 | 736,584 | 332,657 | |
| 正会員会費 | 936,000 | 936,000 | 0 | 117名 (8,000 円×117名) |
| 特別会員 | 60,000 | 60,000 | 0 | 10 名 (6,000 円×10 名) |
| 学生会員会費 | 16,000 | 12,000 | 4,000 | 4名(4,000円×4名) |
| 賛助会員会費 | 130,000 | 140,000 | ▲10,000 | 7社 (10,000円×13口) |
| 学会誌売上金 | 150,000 | 120,000 | 30,000 | |
| 広告料 | 20,000 | 20,000 | 0 | |
| 別刷り代等 | 500,000 | 500,000 | 0 | 別刷り,超過頁,カラー印刷費 |
| 掲載料 | 700,000 | 800,000 | ▲100,000 | 14 編×50,000 円 |
| 雑収入 | 100,000 | 100,000 | 0 | |
| 寄付金 | 0 | 0 | 0 | |
| 収入合計 | 3,681,241 | 3,424,584 | 256,657 | |

| 支出の部 | | | | |
|------------|-----------|-----------|---------|---------------|
| 費目 | 24 年度予算 | 23 年度予算 | 増減 | 備考 |
| 学会誌印刷費 | 1,800,000 | 1,800,000 | 0 | 4 冊×450,000 円 |
| 送料・通信費 | 100,000 | 150,000 | ▲50,000 | |
| 事務費 | 700,000 | 700,000 | 0 | 人件費,事務用品,封筒他 |
| 交通費 | 20,000 | 20,000 | 0 | |
| 会議費 | 15,000 | 5,000 | 10,000 | |
| 学会賞経費 | 50,000 | 50,000 | 0 | 賞状他 |
| 雑費 | 25,000 | 25,000 | 0 | 振込み手数料他 |
| 次年度繰越(予備費) | 971,241 | 674,584 | 296,657 | |
| 支出合計 | 3,681,241 | 3,424,584 | 256,657 | |

日仏海洋学会 役員・評議員(2012-2013年度) 会 長:小松輝久 副会長:森永 勤 吉田次郎 顧 問:今脇資郎 編集委員長:吉田次郎 幹 事: (庶務) 荒川久幸, 長井健容 (会計)神田穰太,市川 香 (編集)田中祐志,磯田 豊 (研究) 北出裕二郎, 千手智晴 (涉外)小池康之,中野俊樹 (広報)内田 裕,柳本大吾 監 事:野村英明, 宮崎奈穂 評議員:荒川久幸,石坂丞二,石丸 隆,磯田 豊, 市川 香, 今脇資郎, 内田 裕, 神田穰太, 北出裕二郎,小池勲夫,小池 隆,河野 博, 小松輝久, 斉藤誠一, 千手智晴, 田中祐志, 中田英昭,長島秀樹,森永 勤,門谷 茂, 柳 哲雄,山口征矢,山崎秀勝,吉田次郎 (選挙選出 24 名) 小池康之, 中野俊樹, 長井健容, 柳本大吾 (会長推薦4名) 以上28名 賞委員: 2012年度:吉田次郎,瀬川 進,千手智晴,有元貴文, 北出裕二郎,小松輝久,荒川久幸,河野 博,石坂丞二 以上9名 2.6月10日(日),日仏会館(東京都恵比寿)におい て2012年度研究発表会を開催した。 日 時: 2012年6月10日(日)10時00分~15時40 分 場 所:日仏会館 501 会議室 プログラムは以下の通り。 10:00-11:20座長 吉田次郎(海洋大) 1. Hypolimnetic turbulence enhancement and superposition of large-scale internal waves in a strongly stratified Lake Biwa (Japan) Guillaume Auger¹, Hidekatsu Yamazaki¹, Hikaru Honma¹, Takeyoshi Nagai1, Michio Kumagai² (1: TUMSAT, 2: Ritsumeikan University) 2. ビンセネス湾沖で係留観測によって捉えた南極底 層水の特性 ○北出裕二郎¹,嶋田啓資²,田村岳史³,深町 康², 青木 茂², 大島慶一郎², 牛尾収輝³ (1:海洋大環境,2:北大低温研,3:極地研) 3. Effect of Turbulent Mixing on Modification of Water Masses in the Central Equatorial Pacific ○ Lingqiao Cheng, Yujiro Kitade (TUMSAT) 4. 日本南岸の黒潮流路変動 前田恵理子¹,出口大貴¹.2,根本雅生¹,○吉田次郎¹ (1:東京海洋大学,2:神奈川海洋科学高校)

- 13:40-15:20 座長 飯淵 敏夫(海生研)
 - 5. Phosphorus stress of microphytoplankton community in the western subtropical North Pacific.
 Mathias GIRAULT, Hisayuki ARAKAWA, Fuminori HASHIHAMA (TUMSAT)
 - 6. 海底堆積粒子の粒径の相違による褐藻類アラメ配 偶体の生残・成長への影響

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○渡辺隼人,松本陽,荒川久幸(海洋大環境)
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- アイゴは海藻と動物のどちらを選択するか?
 ○柴田玲奈¹,片山知史²,荒川久幸³,齊藤 肇¹
 (1:水工研,2:東北大・農,3:海洋大環境)
- 8. 東シナ海における冬季流れ藻の分布と輸送に関す る研究
 - ○水野紫津葉¹,小松輝久¹,鰺坂哲朗²,福田正浩¹, 國分優孝¹,Alabsin natheer¹,阪本真吾¹, 青木優和³
- (1:東大・大海研, 2:京大院・農, 3:東北大・農) 9. 仙台湾におけるアカガイの加入量変動

○佐々修司¹・佐々木浩一・南卓志・片山知史² (1:東大・大海研, 2:東北大・農)

- 15:30-16:10 総会
- 16:20-16:30 2012 年度日仏海洋学論文賞の授賞式

 小野敦史会員(東京海洋大学)
 Distribution and population structure of salps off Adelie Land in the Southern Ocean during austral summer, 2003 and 2005, 48 (2), 55-70, 2010

 柴田玲奈会員(水産総合研究センター) アイゴ成魚に対する動物性餌料の重要性,48(3-4), 103-111,2010

3. 入会

| 氏名 | 所属 |
|-------|---|
| 田上英明 | 海洋政策研究財団 〒105-0001 東京都港区虎ノ門 3-4-10 虎ノ門 35 森ビル |
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| 佐々修司 | 東京大学海洋研究所行動生態計測 〒277-0882 千葉県柏市柏の葉 5-15 |

4. 寄贈図書および資料

農工研ニュース(農村工学研究所); No. 75-79, 神奈川県立博物館研究報告; 41, FRAN NEWS; 28-30, 水産総合研究センター研究報告;35-36, 広島観光コンベンション; Vol. 84-86, Ship & Ocean Newsletter (海洋政策研究財団); No. 265-285, 農村工学研究所報告;51, なつしま (JAMSTEC); 306-315, 水産技術;第4巻1-2, J-STAGE NEWS (独立行政法人科学技術振興機構); No. 29-3, 「海一自然と文化」(東海大学海洋学部); Vol. 9 No. 2, No. 42-45, 高知大学海洋生物研究報告(高知大学総合研究センタ); No. 26, 海洋白書;2012年度, 日本海ブロック資源研究会報告;21-22年度, 養殖研究レター; Vol. 2, Ocean Breeze ; 5–8, 日仏獣医学会誌; Vol. 22, 1-2, ATOMOSPHERE AND OCEAN RESEARCH INSTI-TUTE THE UNIVERSITY OF TOKYO; 2011, 増養殖研究レター;1, 日仏生物学会誌; Vol. 51, ABSCJF; No. 42, 気候システムニュース; No. 2, 中国海洋大学学報; Vol. 41 187-199, 42 200-203, Progress in fishery sciences; Vol. 32 4-6, 33 1-2, Oceanologia et Limnologia sinica; Vol. 41 1-6, Meereswissenchaftliche Berichte Marine Science Repoorts; Vol. 84-85, Annual Report; 2011 年度版, Chinese journal of Oceanology and limnology; Vol. 29 1-5

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日仏海洋学会入会申込書

(正会員・学生会員)

| | | 年度より入会 | | | 年 | 月 | 日申込 |
|---------------------|------------|---------|----|---------|---|---|-----|
| 氏名 | | | | | | | |
| ローマ字 | | | | | 年 | 月 | 日生 |
| 住所 | 听 〒 | | | | | | |
| 勤務先 機関名 | | | | | | | |
| 電 言 | 話 | E-mail: | | | | | |
| 自宅住所 | 听 〒 | | | | | | |
| 電言 | 話 | | | E-mail: | | | |
| 紹介会員氏名 | | | | | | | |
| 送付金額 | | 円 | 送金 | 方法 | | | |
| 会誌の送り先(希望する方に○をつける) | | | 勤 | 務先 | 自 | 宅 | |

(以下は学会事務局用)

| 受付 | 名簿 | 会費 | あて名 | 学会 |
|----|----|----|-----|----|
| | 原簿 | 原簿 | カード | 記事 |

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(財) 日仏会館内

日仏海洋学会

郵便振替番号:00150-7-96503